Height Assessments in Children With Neurofibromatosis Type 1

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Abstract

Previous studies have suggested that children with neurofibromatosis type 1 (NF1) are shorter than their unaffected counterparts. Unfortunately, these reports did not consider other contributing factors that might also influence short stature. The purpose of the current study was to characterize the genetic influence of NF1 on the growth of children. Height data were measured and recorded for 170 patients, whereas parental measurements were obtained for 61 patients to calculate sex-corrected mid-parental target heights. Children with NF1 had population mean height and mid-parental height z-scores statistically different from the general population. Importantly, these differences were pronounced when neither parent had NF1 but were not significant when one of the parents had NF1. Moreover, height z-scores for children with NF1 were also statistically different than their unaffected siblings. Collectively, these data establish a clear effect of a germline NF1 gene mutation on stature in children with NF1.

Keywords

neurofibromatosis type 1, stature, midparental target height, growth

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Neurofibromatosis type 1 (NF1) is an autosomal dominant inherited tumor predisposition syndrome that affects approximately 1 in every 3000 individuals worldwide.¹ In this regard, the defining features of NF1 include tumors of the central and peripheral nervous system (gliomas and neurofibromas). However, children with NF1 are also prone to the development of non-neoplastic features, such as pigmentary abnormalities (cafe-au-lait spots, skinfold freckling, and Lisch nodules), skeletal anomalies, and short stature.¹² Studies examining short stature in children with NF1 have reported heights 2 standard deviations or more below the reference population mean in 13% of affected individuals,³ heights that fell below the 3rd centile for the general population in 15% of children with NF1,⁴ or heights below the 5th centile in 27% to 33% of affected individuals.¹⁵ Although these reports highlight the frequency of short stature in this unique patient population, they fail to account for growth differences between affected children and their parents and unaffected siblings, and do not exclude potential NF1-associated comorbid etiologies for short stature.

To our knowledge, this is the first study to compare growth in children with NF1 to their affected and unaffected family members.

Methods

An unselected, consecutive series of 231 patients with NF1 were seen at a tertiary care, subspecialty referral neurofibromatosis clinic at St Louis Children’s Hospital between May 19, 2010, and October 27, 2011. The diagnosis of NF1 was made according to established diagnostic criteria by a single NF specialist (DHG).²⁶ The diagnosis of NF1 was made solely by clinical findings, and no genetic studies were performed.

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Patients were recruited to participate in the study if they were more than 3 years of age and were able to stand independently for a height measurement. Patients were excluded if they had a known risk for disordered growth: precocious puberty or other endocrine disorder, intracranial tumor, skeletal abnormality, or treatment with attention-deficit hyperactivity disorder (ADHD) medication. No patients with skeletal anomalies (scoliosis, pseudarthrosis, long bone deformities, etc) were included in order to exclude any such effect on patient height. For eligible subjects with NF1, parents were invited to participate in the study unless they had a history of a specific growth disorder in childhood or had a known skeletal disorder. Siblings were invited to participate unless they had a significant medical illness, a known endocrine disorder, or a skeletal abnormality. One hundred seventy primary subjects agreed to participate in the study (94 females and 76 males). Patient demographics, including sex, age, and pertinent clinical information, was also recorded.

Of the 170 patients, parental height information on both parents was measured and recorded for 61 patients, and sibling height information was measured and recorded for 19 patients. Sex-corrected mid–parental target height was calculated by taking the arithmetic mean of parental height and adding or subtracting 6.5 cm for males or females, respectively. This mid–parental target height was used to estimate the genetically determined adult height percentile that each patient should obtain if they were unaffected by NF1. Parental and sibling NF1 status (affected or unaffected) was noted.

Patients’ heights, midparental target height values, and siblings’ heights were standardized to obtain z scores using the Centers for Disease Control and Prevention (CDC) 2000 growth charts for reference population norms of stature by age and sex. The z scores were summarized using means and standard deviations. For patients older than age 20 years and for the mid–parental target height z score calculation, we used the z score corresponding to a 20-year-old, which was the upper age limit of the CDC population growth charts. Because the average onset of puberty is age 10 years for girls and age 12 years for boys, patient z scores for girls 10 years and younger were compared to those older than age 10 years, and for boys 12 years and younger, they were compared to those older than age 12 years to assess the effect of NF1 on puberty.

One-sample t tests were used to see whether the average of patients’ z scores (overall and grouped by age) were significantly different from the population level, whereas paired t tests were used to compare the average z scores of patients versus parents (midparental target height) as well as patients versus unaffected siblings. All analyses were 2-sided, and significance was set as P value < .05. Statistical analyses were performed using the SAS statistical package (SAS Institutes, Cary, North Carolina).

The study protocol was approved by the Human Research Protection Office at the Washington University School of Medicine.

Results

Heights for children with NF1 were plotted on CDC growth charts (Figures 1 and 2 for boys and girls, respectively) and represented as height z score data (Table 1). The frequency distribution of standardized z scores relative to the population mean height revealed that 116 (68%) children had negative z scores relative to the population mean height (Figure 3). These affected individuals had a mean height z score of –0.50 ± 1.27, which was significantly shorter than that observed in the general population (P < .001).

The frequency distribution of z score values relative to midparental target height revealed a calculated mid–parental target height z score of 0.06 ± 0.86 (Figure 4). The affected patient height z scores were significantly different from the mid–parental target height z scores (P < .001). This difference was pronounced when neither parent has NF1 (P < .01) but was not significant when one of the parents had NF1 (P = .1). As predicted, the unaffected sibling z scores differed from those seen in affected children (P = .01) but were not different from the mid–parental...
of short stature, the impact of a germline \textit{NF1} gene mutation on growth, and the contribution of \textit{NF1}-associated comorbid features. In an effort to address these issues, we used standardized height measurements routinely employed by endocrinologists consulted for the evaluation of short stature.

First, using a combination of population mean height and mid–parental target height assessments, we estimate that true short stature in this affected population, defined as >2 standard deviations below the predicted mean, is observed in 8\% of children with \textit{NF1}. This estimate of population frequency is lower than that reported by Szudek\textsuperscript{3} (13\%) and Clementi\textsuperscript{4} (15\%); however, these studies did not exclude subjects with scoliosis, pseudarthrosis, or other primary causes for short stature. Similar to learning disabilities in children with \textit{NF1},\textsuperscript{11,12} the distribution of stature in affected individuals was left shifted, supporting the conclusion that stature is reduced to some degree in all individuals with \textit{NF1}, even if they do not have frank short stature.

Second, standardization for age and sex by \textit{z} score allowed us to combine measurements across groups that differed in age and sex. Using this method, we showed that short stature reflects the presence of a germline \textit{NF1} mutation in bone and somatic growth. In these studies, genetically engineered mouse strains lacking \textit{Nf1} expression in bone have skeletal abnormalities,\textsuperscript{16,17} whereas genetically engineered mouse strains lacking \textit{Nf1} expression in bone have skeletal abnormalities,
expression in central nervous system neurons have reduced hypothalamic-pituitary axis function and lower insulin-like growth factor–I (IGF-1) levels. Future work in both preclinical models and in children with NF1 will be required to determine how these factors contribute to short stature in the pediatric population.

Third, short stature could reflect comorbid NF1-associated features, such as precocious puberty, hypothalamic/optic chiasmal gliomas, treatment with stimulant medication, or skeletal abnormalities (pseudarthrosis, scoliosis). In our series, we specifically excluded individuals with NF1 and short stature because of these causes. Precocious puberty can lead to accelerated bone maturation and result in short stature from premature epiphyseal closure. Our population did not show a difference in height z scores when we compared prepubertal (girls younger than age 10 years, boys younger than age 12 years) to postpubertal-aged children, so short stature did not appear to be due to early timing of puberty, which would lead to taller stature in early childhood and shorter stature after puberty. Suprasellar lesions, including hypothalamic glioma, may cause growth hormone deficiency or precocious puberty. ADHD is also prevalent in children with NF1 compared with the healthy population, and stimulant treatment has been associated with growth deficits. Finally, growth hormone deficiency in the absence of a suprasellar lesion may be another cause for short stature.

Future studies on individuals with NF1 should focus on the potential mechanisms, leading to short stature in this at-risk population. In this regard, characterization of longitudinal growth patterns will be necessary to determine whether pubertal growth is affected more than prepubertal growth. Similarly, an evaluation of circulating growth-related factors (growth hormone, IGF-1, and IGF-binding proteins) may reveal abnormalities in growth hormone secretion or action. Lastly, measurements of the responses of bone to usual growth mediators may show differences in bone at the cellular level. Collectively, these investigations may further elucidate the causes of NF1-mediated growth dysfunction and are likely to define the molecular and cellular etiologies for growth failure in children with NF1 currently classified as having idiopathic short stature.

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Author Contributions
EAS, DVO, and NLN collected clinical data. EAS and FG performed statistical analysis. EAS, DHG, and ASH prepared the manuscript.

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Ethical Approval
The study protocol was approved by the Human Research Protection Office at the Washington University School of Medicine.

References